

Evaluation of multiple sclerosis disability outcome measures using pooled clinical trial data

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Abstract

Objective

We report analyses of a pooled database by the Multiple Sclerosis Outcome Assessments Consortium to evaluate 4 proposed components of a multidimensional test battery.

Methods

Standardized data on 12,776 participants, comprising demographics, multiple sclerosis disease characteristics, Expanded Disability Status Scale (EDSS) score, performance measures, and Short Form–36 Physical Component Summary (SF-36 PCS), were pooled from control and treatment arms of 14 clinical trials. Analyses of Timed 25-Foot Walk (T25FW), 9-Hole Peg Test (9HPT), Low Contrast Letter Acuity (LCLA), and Symbol Digit Modalities Test (SDMT) included measurement properties; construct, convergent, and known group validity; and longitudinal performance of the measures individually and when combined into a multidimensional test battery relative to the EDSS and SF-36 to determine sensitivity and clinical meaningfulness.

Results

The performance measures had excellent test–retest reliability and showed expected differences between subgroups based on disease duration and EDSS level. Progression rates in detecting time to 3-month confirmed worsening were lower for T25FW and 9HPT compared to EDSS, while progression rates for LCLA and SDMT were similar to EDSS. When the 4 measures were analyzed as a multidimensional measure rather than as individual measures, progression on any one performance measure was more sensitive than the EDSS. Worsening on the performance measures analyzed individually or as a multidimensional test battery was associated with clinically meaningful SF-36 PCS score worsening, supporting clinical meaningfulness of designated performance test score worsening.

Conclusion

These results support the use of the 4 proposed performance measures, individually or combined into a multidimensional test battery as study outcome measures.

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Glossary

9HPT = 9-Hole Peg Test; **BDI** = Beck Depression Inventory; **CI** = confidence interval; **EDSS** = Expanded Disability Status Scale; **HRQoL** = health-related quality of life; **ICC** = intraclass correlation coefficient; **LCLA** = Low Contrast Letter Acuity; **MCS** = Mental Component Summary; **MS** = multiple sclerosis; **MSFC** = Multiple Sclerosis Functional Composite; **MSOAC** = Multiple Sclerosis Outcome Assessments Consortium; **PASAT** = Paced Auditory Serial Addition Test; **PCS** = Physical Component Summary; **RR** = relapsing-remitting; **SDMT** = Symbol Digit Modalities Test; **SF-36** = Short Form-36; **T25FW** = Timed 25-Foot Walk.

Recognition of the limitations of existing measures of multiple sclerosis (MS)-related disability¹ led to development of the Multiple Sclerosis Functional Composite (MSFC) as an alternative clinical outcome measure.^{2,3} The MSFC integrated 3 quantitative performance measures: Timed 25-Foot Walk (T25FW), 9-Hole Peg Test (9HPT), and Paced Auditory Serial Addition Test (PASAT).^{2,3} Component test scores were normalized by conversion to *z* scores using a reference sample and averaged to create the MSFC score. Subsequent utilization of the MSFC in clinical trials and other studies highlighted several advantages: high reliability; coverage of cognition, which is not adequately captured by other outcome measures; a single score on a continuous scale; ease of use; and capacity to be administered by a trained technician rather than a physician. However, while utilization of *z* scores offered statistical advantages, it impeded MSFC acceptance and adoption. Both regulatory agencies and clinicians had difficulty interpreting the clinical meaningfulness of the MSFC score.

Subsequent research provided guidance around thresholds of MSFC measure variability and clinically meaningful change.⁴⁻⁷ In addition, researchers recommended adding Low Contrast Letter Acuity (LCLA), a test of visual impairment validated in MS,^{8,9} and replacing the PASAT with the Symbol Digit Modalities Test (SDMT).¹⁰ In this article, we report the initial analyses of a pooled database to assess measurement properties; sensitivity; construct, convergent, and known group validity; and clinical meaningfulness of 4 performance tests (T25FW, 9HPT, LCLA, and SDMT) individually and combined into a multidimensional outcome measure of MS-related disability.

Methods

Study design

The Multiple Sclerosis Outcome Assessments Consortium (MSOAC) was established in 2012 to accelerate development and validation of improved clinical outcome measures of MS-related disability.^{11,12} The organization of MSOAC and overall approach to develop and validate clinical outcome measures of MS-related disability have been reported previously.¹³ Frequent interactions with the European Medicines Agency and US Food and Drug Administration contributed to MSOAC's research approach.¹³ After executing data use agreements, MSOAC obtained prospectively acquired patient-level data

from 16 clinical trials including 14,370 participants and combined these into a single database. Fourteen of these trials, comprising 12,776 patients, were used for the analyses reported here.

To allow aggregation of data from distinct datasets, a common data standard was developed through the Coalition for Accelerating Standards and Therapies, and data from each clinical trial were remapped to the Clinical Data Interchange Standards Consortium data standard to create a single pooled data set.¹⁴ The standardized data comprising the control and treatment arms of clinical trials formed the MSOAC database. Data elements included demographics, MS disease characteristics, treatment, relapse data, Expanded Disability Status Scale (EDSS) score, performance measures, and patient-reported outcomes.¹³ The MSOAC database does not include the MRI data.

Outcome measures

These analyses focused on 4 performance measures and their relation to other measures in the database: T25FW (short distance walking speed to measure ambulation),¹⁵ 9HPT (dexterity and upper extremity motor function),¹⁶ LCLA (vision),⁹ and SDMT (cognitive processing speed and sustained attention).¹⁰ Data on PASAT are presented elsewhere.¹⁷ The EDSS is an ordinal scale ranging from 0 to 10 based on the severity of findings on the neurologic examination, walking ability, and ability to carry out activities of daily living, with higher scores indicating worse disability.¹⁸ The Short Form-36 (SF-36) is a 36-item questionnaire that includes 8 multi-item health concepts (Physical Functioning, Role-Physical, Bodily Pain, General Health, Vitality, Social Functioning, Role-Emotional, and Mental Health).¹⁹ Scores are a mean of subsetted questions and range from 0 to 100; higher scores indicate better health-related quality of life (HRQoL). The SF-36 has 2 summary scales, the Physical Component Summary (PCS) and the Mental Component Summary (MCS), whose calculation produces a T-score, with a mean score of 50 and SD of 10, representing the reference score for the US general population. The Beck Depression Inventory (BDI) is a 21-item self-report measure of depression with scores ranging from 0 to 62 and higher score indicating more severe depression symptoms.²⁰

For these analyses, worsening was defined as follows: T25FW (20% increase),¹⁵ 9HPT (20% increase),¹⁶ LCLA with 2.5% contrast (20% or 7-letter decrease),⁹ SDMT (4-point

decrease),¹⁰ EDSS (baseline score 0: 1.5-point increase, baseline score 1.0–5.5: 1-point increase, baseline score ≥ 6.0 : 0.5-point increase),²¹ and SF-36 PCS Score (5-point worsening).²² For all measures except SF-36 PCS, the worsening had to be sustained for at least 3 months.

Statistical methods

No imputation was done for missing data other than for participants unable to complete the T25FW or 9HPT because of disability. Following convention, imputation for patients unable to perform was 180 seconds for T25FW and 300 seconds for 9HPT.²³ The MSFC administration and scoring manual states that for T25FW testing patients should use their usual assistive devices and an effort should be made to use the same device over the course of the study. Summary scores of the SF-36 MCS and PCS were calculated using standard methods, which provide T-scores for analysis. For the SF-36 8 health concept scores, Quality Metrics Health Outcomes Scoring Software was utilized. The maximum data recovery method was used to handle missing data. If any individual item was missing for the BDI score, the total score was not calculated for that participant and time point.

Test–retest reliability was assessed by intraclass correlation coefficient (ICC) of all administrations of each test (2–6 compared with test 1) based on periods in which patient status on the EDSS did not change and not exceeding 6 months from baseline.²⁴ Correlations among the EDSS and performance tests were assessed by Spearman rank correlation coefficient. Time to confirmed clinically meaningful worsening was analyzed by Kaplan–Meier methods. Cohen kappa coefficient was used to assess agreement in worsening in different disability measures. Baseline outcome measure scores were compared. The baseline score for each performance measure was compared between the groups of patients based on disease duration and EDSS score using an analysis of variance model adjusting for age in 5-year age bands.

Data availability

Per data use agreements, analyses were done through a contract research organization (Premier Research) under the oversight of the Critical Path Institute. Pooled data were not available to individual sponsors or academic members, although per agreement, placebo data were made available publicly.²⁵ Consortium members contributed to development of the statistical analysis plan, and had access to results from all analyses. The authors had full access to all the data generated in this fashion.

Results

Table 1 summarizes the data available, baseline demographics, disease characteristics, EDSS score, performance test results, and participant self-reported measures. Overall, the population was relatively young with a recent diagnosis of MS, predominantly relapsing-remitting (RR) course, and mild

disability. Although fewer studies included LCLA, SDMT, and self-reported measures, substantial data were available for all outcome measures.

The frequency distributions of the T25FW and 9HPT were positively skewed and showed floor effects, with scores tending to be clustered at shorter times (figure 1). Both possessed the ability to distinguish gradations of performance in the middle of the scale. LCLA distribution appeared mildly negatively skewed without floor or ceiling effects. SDMT scores showed no evidence of skewing, or floor or ceiling effects.

Table 2 summarizes trends over the first 6 assessments for the performance tests. T25FW, 9HPT, and LCLA tended to worsen over time and showed minimal or no practice effects, while the SDMT demonstrated modest practice effects. Test–retest reliability was estimated by calculating the ICC, accounting for practice effects where needed. All of the measures showed good test–retest reliability, though the ICC for T25FW was somewhat lower (0.71) compared to the other tests (0.84–0.88).

To compare sensitivity to change of the performance measures with EDSS, time from baseline to 3-month confirmed worsening over 24 months was analyzed (figure 2). The study populations available for each comparison differed, leading to differing proportions with 3-month confirmed worsening on EDSS. Using a 20% threshold for T25FW, 6.5% worsened compared to 20.2% on EDSS. Using a 20% threshold for 9HPT, 2.9% worsened compared to 20.2% on EDSS. Using 7-point threshold for LCLA, 13.1% worsened compared to 16.1% on EDSS. Using 4-point threshold for SDMT, 15.0% worsened at 18 months compared to 11.4% on EDSS at 18 months and 14.5% at 24 months. Thus, progression rates were lower for T25FW and 9HPT compared to that of EDSS, while progression rates for LCLA and SDMT were similar to or higher than compared to EDSS. When the performance tests were combined into a multidimensional outcome measure, the proportion of participants worsening on any one performance test was substantially greater than the proportion worsening on EDSS. When worsening on 2 performance tests was required, sensitivity to disability progression was somewhat reduced compared to the EDSS. The progression events defined by the performance tests were weakly associated with or independent of those defined by the EDSS: T25FW (Cohen $\kappa = 0.02$, 95% confidence interval [CI] –0.00 to 0.03), 9HPT ($\kappa = 0.00$, 95% CI –0.01 to 0.01), LCLA ($\kappa = 0.11$, 95% CI 0.08 to 0.14), and SDMT ($\kappa = -0.02$, 95% CI –0.06 to 0.02).

To investigate construct and convergent validity, correlations between the performance measures and EDSS were analyzed (table 3). The T25FW and 9HPT correlated strongly with one another and demonstrated the strongest correlation to the EDSS relative to other performance measures. LCLA and SDMT were weakly correlated to the other performance

Table 1 Baseline characteristics

Measure	N	Mean (SD) or n (%)	Median (range)
Age, y	12,776	39.5 (9.92)	40.0 (17–72)
Sex, n (%)	12,776		
Female		8,799 (68.9)	
Male		3,977 (31.1)	
Disease course	12,776		
Relapsing-remitting		10,789 (84.4)	
Secondary progressive		1,044 (8.2)	
Primary progressive		943 (7.4)	
Disease duration, y	6,641	6.5 (7.26)	4.0 (0–48)
EDSS	12,776	2.9 (1.63)	2.5 (0–8.0)
T25FW, s	11,649	7.6 (9.84)	5.4 (1.0–231.0) ^a
9HPT, s	11,653	24.3 (14.30)	21.3 (5.0–331.0) ^a
LCLA 2.5% contrast, number correct	5,669	34.6 (11.65)	37.0 (0–60)
SDMT, number correct	2,583	47.9 (15.90)	48.0 (0–110)
PCS	7,766	41.5 (9.95)	40.9 (10–73)
MCS	7,766	47.7 (11.53)	49.5 (–5 to 74)
BDI	2,824	8.9 (8.62)	7.0 (0–53)

Abbreviations: 9HPT = 9-Hole Peg Test; BDI = Beck Depression Inventory; EDSS = Expanded Disability Status Scale; LCLA = Low Contrast Letter Acuity with 2.5% contrast; MCS = Short Form–36 Mental Component Summary; PCS = Short Form–36 Physical Component Summary; SDMT = Symbol Digit Modalities Test; T25FW = Timed 25-Foot Walk.

^a The extreme values were verified as being present in the database.

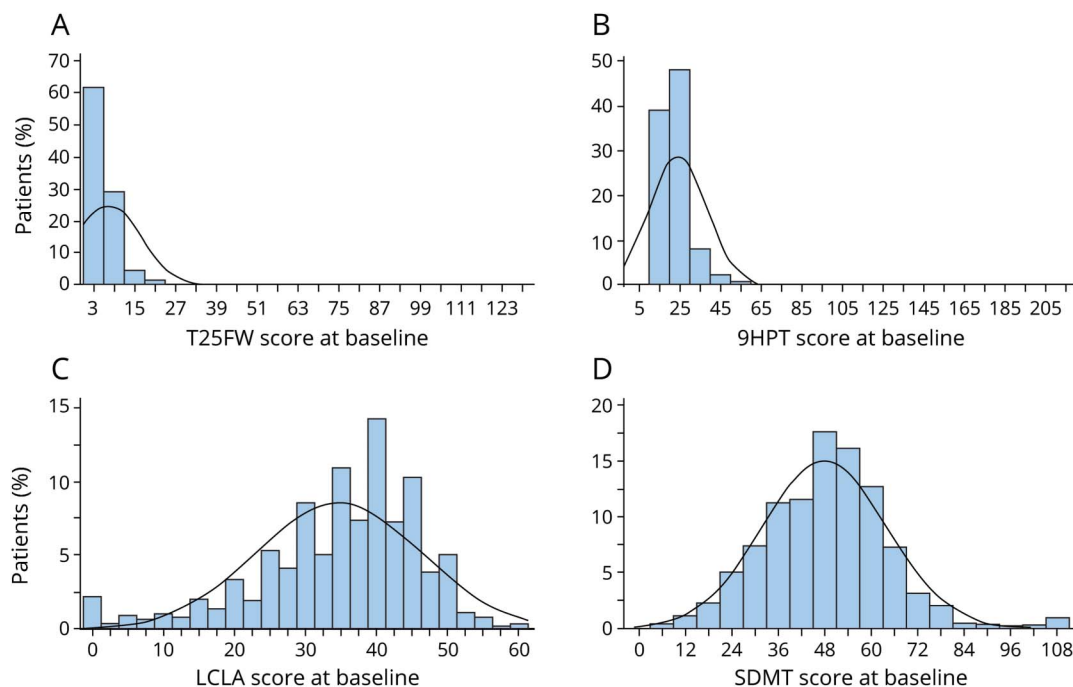
measures and EDSS. Between the two, the SDMT had somewhat stronger correlation to the EDSS. Cross-sectional correlations among outcomes at baseline were notably stronger than the correlations among changes from baseline to endpoint, which had a similar pattern of correlative strength (T25FW > 9HPT > SDMT > LCLA), but wholly weaker in magnitude.

Known group validity was assessed as a function of disease duration and disability level (table 4). At baseline, values for all 4 performance measures were better in participants with MS of shorter duration (<10 years since symptom onset) compared to those with disease of longer duration (≥10 years). Similarly, the results on all 4 performance tests were better in participants with lower EDSS scores (0–3.5) vs those with higher EDSS scores (4.0–10).

To explore clinical meaningfulness, correlations were calculated between performance measures and participant self-reported measures of HRQoL and depression (table 3). At baseline, T25FW, 9HPT, and SDMT correlated moderately with SF-36 PCS and significantly but weakly with MCS and BDI. LCLA correlated weakly with SF-36 PCS and MSC, and BDI. Correlations between change baseline to endpoint

in the performance measures and change on SF-36 PCS or MCS, or BDI were generally not significant and weak at best (table 3). Among participants with worsening from baseline to endpoint on the T25FW, 9HPT, or SDMT, the mean SF-36 PCS also worsened ($p < 0.001$, $p < 0.001$, and $p = 0.0308$, respectively) (table 5). Similarly, among participants who showed baseline to endpoint worsening on the T25FW, 9HPT, or SDMT, the proportions of participants with 5-point PCS worsening on SF-36 PCS were greater. Non-significant trends were seen for mean PCS change and the proportion with 5-point PCS change among participants who did or did not experience baseline to endpoint worsening on LCLA. Mean PCS worsened among participants with baseline to endpoint worsening in each of 2 groups: those with worsening on any one measure and those worsening on 2 or more performance measures. The SF-36 PCS was improved or stable, respectively, among participants who did not worsen on 1 or on 2 or more performance measures. Similarly, the proportions of participants with 5-point SF-36 PCS worsening were greater among participants who showed baseline to endpoint worsening on 1 or on 2 or more performance measures, suggesting that worsening on any single performance measure was clinically relevant.

Figure 1 Distribution of performance measure scores at baseline



(A) Timed 25-Foot Walk (T25FW) (seconds). (B) 9-Hole Peg Test (9HPT) (seconds). (C) Low Contrast Letter Acuity (LCLA) with 2.5% contrast (number correct). (D) Symbol Digit Modalities Test (SDMT) (number correct).

Discussion

We present analyses to characterize the measurement properties; sensitivity; construct, convergent, and known group validity; and clinical meaningfulness of 4 performance measures—T25FW, 9HPT, LCLA, and SDMT—to permit use individually or combined into a multidimensional test battery as a primary or co-primary outcome measure. We have assessed the components of the proposed multidimensional test battery in relation both to the EDSS and self-reported measures of health-related quality of life and depression. These results, based on a database of 14 datasets comprising 12,776 participants, represent the largest pooled analysis of prospectively acquired clinical trial data in MS to date. The demographics of the pooled dataset largely reflect the type of

patients historically enrolled in MS clinical trials, for which trials in RRMS have predominated.

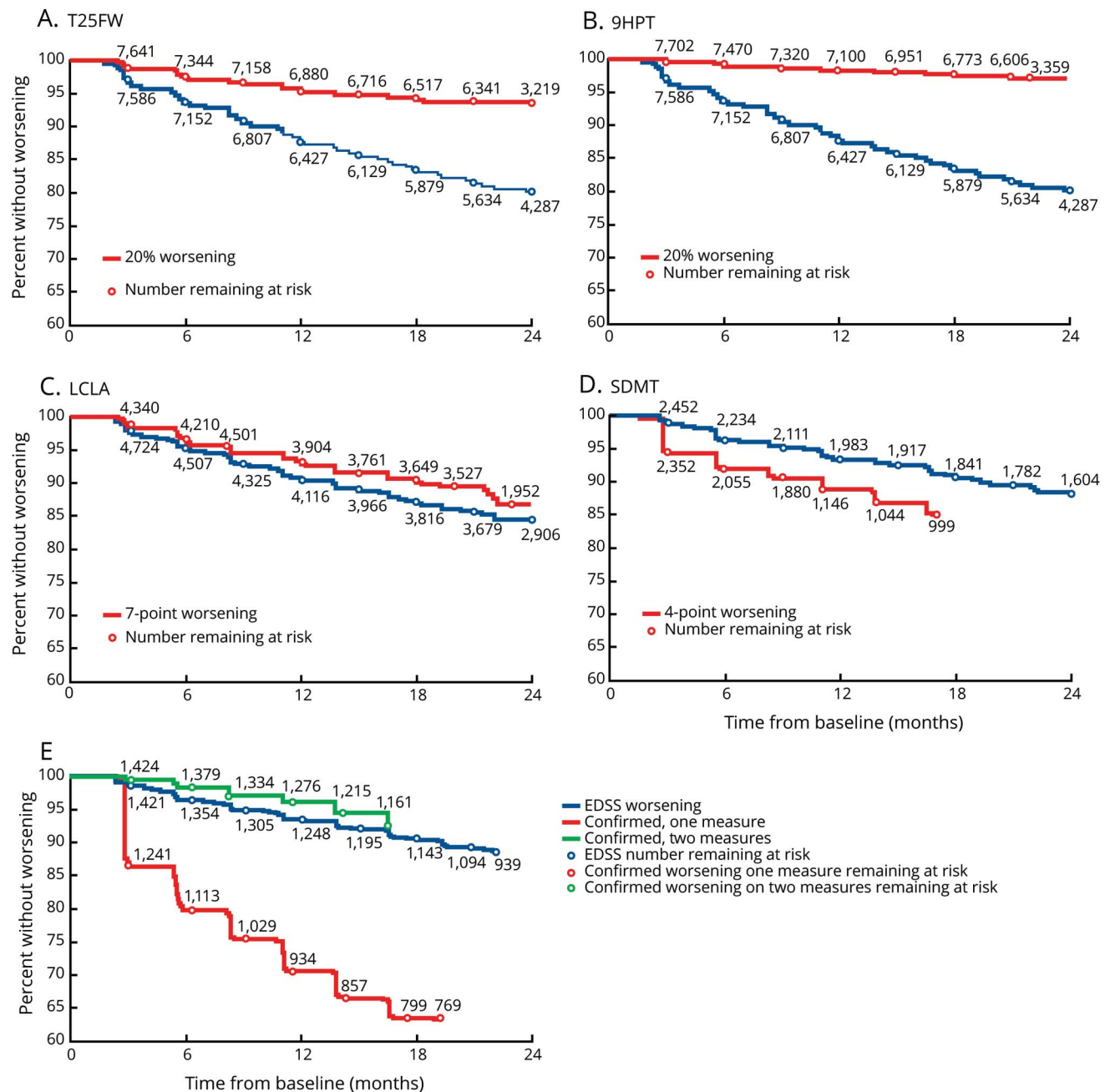
The distributions of the T25FW and 9HPT were positively skewed, and both measures demonstrated floor effects. The potential for a high proportion of patients to perform these measure as well as can be performed by a healthy control can result in reduced ability to distinguish gradations of performance at the lower end of the scale (demonstrated by far left peaks in figure 1, A and B). Baseline LCLA and SDMT scores were more normally distributed, without evidence of floor or ceiling effects. The T25FW, 9HPT, and LCLA showed no clearcut evidence of practice effects. As is typical of most cognitive measures, the SDMT exhibited some practice effects, but these appeared not to affect the normality of the

Table 2 Practice effects and test-retest reliability of performance measures with tests 2–6 each compared to test 1

Test	N	Test 2	Test 3	Test 4	Test 5	Test 6	ICC
T25FW	7,971	0.08	0.08	0.05	0.08	0.13	0.71
9HPT	7,973	0.02	0.00	-0.03	-0.04	0.00	0.84
LCLA	4,611	-0.02	-0.03	-0.01	0.00	0.00	0.88
SDMT	2,094	0.03	0.10	0.15	0.28	0.37	0.85

Abbreviations: 9HPT = 9-Hole Peg Test; ICC = intraclass correlation coefficient (a measure of reliability, higher is better, 1 is the maximum possible score); LCLA = Low Contrast Letter Acuity with 2.5% contrast; SDMT = Symbol Digit Modalities Test; T25FW = Timed 25-Foot Walk. The values for tests 2–6 are the regression coefficients for the 2nd to 6th test, expressed as an effect size to make them comparable. For example, with T25FW, the 2nd test was on average 0.08 SDs higher than the first test.

Figure 2 Kaplan-Meier graphs of time to 3-month confirmed disability worsening on performance measures compared to Expanded Disability Status Scale



(A) Timed 25-Foot Walk (T25FW). (B) 9-Hole Peg Test (9HPT). (C) Low Contrast Letter Acuity (LCLA) with 2.5% contrast. (D) Symbol Digit Modalities Test (SDMT). (E) Any 1 or any 2 performance measures.

SDMT's frequency distribution. All 4 performance measures demonstrated good test-retest reliability, indicating they yield reproducible scores if there is no change in the participant's condition. As a result, changes in a score can be assumed to be due to the participant's condition rather than measurement variability. These results support the advantageous measurement properties of the 4 performance measures.

These results provide a cautionary note regarding the population for which these measures will be most useful. The

majority of participants represented in the pooled dataset had RRMS with relatively mild disability. In turn, the T25FW and 9HPT exhibited floor effects, which may explain the decreased sensitivity of 3-month confirmed worsening of T25FW and 9HPT compared to EDSS. More sensitive tests may be needed in studies enrolling participants with mild gait and upper extremity impairments.²¹ One might question whether confirmed worsening on EDSS at the low end of the scale in patients with MS with mild impairment represents increasing disability, or simply new signs on the neurologic examination.

Table 3 Correlations between outcome measures

	9HPT	LCLA	SDMT	EDSS	PCS	MCS	BDI
Baseline correlations							
T25FW	0.52 (0.51 to 0.53)	-0.30 (-0.32 to -0.27)	-0.42 (-0.46 to -0.38)	0.56 (0.55 to 0.58)	-0.40 (-0.42 to -0.38)	-0.13 (-0.16 to -0.11)	0.22 (0.18 to 0.26)
9HPT		-0.33 (-0.35 to -0.31)	-0.47 (-0.51 to -0.43)	0.54 (0.53 to 0.56)	-0.33 (-0.36 to -0.31)	-0.14 (-0.16 to -0.11)	0.20 (0.16 to 0.24)
LCLA			0.34 (0.30 to 0.39)	-0.29 (-0.31 to -0.27)	0.12 (0.09 to 0.14)	0.19 (0.16 to 0.22)	-0.16 (-0.20 to -0.12)
SDMT				-0.34 (-0.38 to -0.29)	0.36 (0.32 to 0.41)	0.21 (0.16 to 0.26)	-0.20 (-0.24 to -0.15)
	9HPT change	LCLA change	SDMT change	EDSS change	PCS change	MCS change	BDI change
Correlations of change from baseline to endpoint							
T25FW change	0.30 (0.28 to 0.32)	-0.08 (-0.11 to -0.06)	-0.14 (-0.19 to -0.09)	0.29 (0.27 to 0.31)	-0.20 (-0.23 to -0.18)	-0.09 (-0.12 to -0.06)	0.10 (0.05 to 0.14)
9HPT change		-0.06 (-0.09 to -0.04)	-0.20 (-0.25 to -0.15)	0.23 (0.22 to 0.25)	-0.16 (-0.19 to -0.13)	-0.07 (-0.10 to -0.05)	0.11 (0.07 to 0.16)
LCLA change			0.06 (0.01 to 0.11)	-0.11 (-0.13 to -0.08)	0.02 (-0.01 to 0.05)	0.06 (0.03 to 0.10)	-0.02 (-0.07 to 0.03)
SDMT change				-0.12 (-0.16 to -0.08)	0.00 (-0.01 to 0.05)	0.06 (0.03 to 0.10)	-0.09 (-0.13 to -0.04)

Abbreviations: 9HPT = 9-Hole Peg Test; CI = confidence interval; LCLA = Low Contrast letter Acuity with 2.5% contrast; SDMT = Symbol Digit Modalities Test; T25FW = Timed 25-Foot Walk. Values are Spearman correlation coefficients (95% CI).

At baseline, the T25FW and 9HPT had stronger correlations with the EDSS and with each other than with the other performance measures. These results support the construct validity of the T25FW and 9HPT, as both are measures of physical functions that overlap with the EDSS in its lower range (EDSS 0–4.0) as seen in this population. In comparison, LCLA and SDMT correlated less strongly with EDSS and the other performance measures, supporting their additive value, to assess functions not captured by the other performance measures and EDSS. Compared to correlations at baseline, all the correlations for change from baseline to endpoint were much weaker. Cohen kappa coefficients

showed that the confirmed worsening events defined by the 4 performance measures were largely independent of those defined by EDSS. Taken together, these results suggest that the 4 performance measures assess overlapping but somewhat different aspects of disability and disability worsening than does the EDSS.

All 4 performance measures were worse in participants with longer MS disease duration and with worse disability measured by EDSS, supporting known group validity. Exploratory analyses were undertaken to assess the clinical meaningfulness of worsening on the performance measures using the SF-36

Table 4 Known group analysis of baseline values based on disease duration and disability level

	Disease duration, y			EDSS		
	<10	≥10	Difference (95% CI), p value	0–3.5	4.0–10	Difference (95% CI), p Value
T25FW, s	7.7	13.3	N = 5,597, 5.57 (4.74 to 6.40), $p < 0.0001$	6.1	12.7	N = 11,595, 6.63 (6.21 to 7.06), $p < 0.0001$
9HPT, s	24.3	29.9	N = 5,599, 5.57 (4.48 to 6.65), $p < 0.0001$	21.7	31.8	N = 11,594, 10.10 (9.48 to 10.72), $p < 0.0001$
LCLA, number correct	33.2	30.7	N = 3,579, -2.50 (-3.75 to -1.25), $p < 0.0001$	34.8	27.4	N = 5,787, -7.46 (-8.33 to -6.60), $p < 0.0001$
SDMT, number correct	48.5	45.2	N = 2,543, -3.31 (-4.85 to -1.77), $p < 0.0001$	49.8	41.2	N = 2,583, -8.60 (-10.09 to -7.12), $p < 0.0001$

Abbreviations: 9HPT = 9-Hole Peg Test; CI = confidence interval; EDSS = Expanded Disability Status Scale; LCLA = Low Contrast Letter Acuity with 2.5% contrast; SDMT = Symbol Digit Modalities Test; T25FW = Timed 25-Foot Walk.

Table 5 Change in Short Form–36 Physical Component Summary (PCS) in participants with and without worsening on Expanded Disability Status Scale (EDSS) and performance measures

Disability measure	N	Absolute change in PCS (SD) among participants with disability measure worsening	Absolute change in PCS (SD) among participants without disability measure worsening	p Value	Percent (95% CI) with 5-point PCS worsening among participants with disability measure worsening	Percent (95% CI) with 5-point PCS worsening among participants without disability measure worsening	Odds ratio (95% CI), p Value
EDSS	Total: 7,455	–2.75 (8.21)	0.43 (7.54)	$p < 0.0001$	36.6 (34.1 to 39.1)	20.5 (19.5 to 21.6)	2.23 (1.98 to 2.53), $p < 0.0001$
	Worse: 1,479						
	Not worse: 5,976						
T25FW (20%)	Total: 7,455	–2.18 (7.83)	0.37 (7.67)	$p < 0.0001$	33.4 (31.1 to 35.7)	20.9 (19.9 to 22.0)	1.89 (1.68 to 2.13), $p < 0.0001$
	Worse: 1,666						
	Not worse: 5,789						
9HPT (20%)	Total: 7,455	–2.86 (8.33)	0.04 (7.68)	$p < 0.0001$	38.6 (34.7 to 42.5)	22.3 (21.4 to 23.4)	2.18 (1.84 to 2.59), $p < 0.0001$
	Worse: 622						
	Not worse: 6,833						
LCLA (7 point)	Total: 4,678	0.03 (7.95)	0.38 (7.49)	$p = 0.2907$	22.1 (18.8 to 25.7)	20.0 (18.8 to 21.3)	1.13 (0.92 to 1.40), $p = 0.2662$
	Worse: 570						
	Not worse: 4,108						
SDMT (4-point)	Total: 1,467	–1.15 (8.19)	–0.04 (7.69)	$p = 0.0308$	28.8 (23.7 to 34.4)	22.2 (19.9 to 24.7)	1.42 (1.06 to 1.89), $p = 0.0201$
	Worse: 288						
	Not worse: 1,179						
Worse on any 1 performance measure (T25FW or 9HPT or LCLA or SDMT)	Total: 7,455	–1.63 (7.94)	0.51 (7.60)	$p < 0.0001$	30.8 (29.0 to 32.7)	20.2 (19.0 to 21.3)	1.77 (1.58 to 1.97), $p < 0.0001$
	Worse: 2,478						
	Not Worse: 4,977						

Continued

Table 5 Change in Short Form-36 Physical Component Summary (PCS) in participants with and without worsening on Expanded Disability Status Scale (EDSS) and performance measures (continued)

Disability measure	N	Absolute change in PCS (SD) among participants with disability measure worsening	Absolute change in PCS (SD) among participants without disability measure worsening	p Value	Percent (95% CI) with 5-point PCS worsening among participants with disability measure worsening	Percent (95% CI) with 5-point PCS worsening among participants without disability measure worsening	Odds ratio (95% CI), p Value
Worse on any 2 or more performance measures	Total: 7,455	-2.43 (8.26)	-0.00 (7.71)	$p < 0.0001$	22.6 (21.7 to 23.7)	35.4 (31.6 to 39.3)	1.87 (1.57 to 2.23), $p < 0.0001$
	Worse: 616						
	Not Worse: 6,839						

Abbreviations: 9HPT = 9-Hole Peg Test; CI = confidence interval; LCLA = Low Contrast Letter Acuity with 2.5% contrast; SDMT = Symbol Digit Modalities Test; T25FW = Timed 25-Foot Walk.

PCS as an anchor. SF-36 PCS correlated moderately at baseline with T25FW, 9HPT, and SDMT and weakly with LCLA. SF-36 MCS and BDI correlated weakly with all 4 performance measures at baseline. Group aggregate changes from baseline to endpoint in the performance measures and self-report measures correlated weakly or not at all when directionality was not considered. Importantly, for participants experiencing confirmed worsening from baseline to endpoint on the T25FW, 9HPT, and SDMT, the SF-36 PCS was significantly worse, and the likelihood of a 5-point worsening on SF-36 PCS, considered clinically important,²² was significantly greater. The LCLA results mirrored these findings but with statistically nonsignificant trends. These results indicate that the T25FW, 9HPT, and SDMT assess clinically meaningful aspects of MS-related disability and that the proposed thresholds for clinically meaningful change for each are reasonable. The nonsignificant trend of concomitant worsening in the LCLA and SF-36 PCS provides some support for the clinical meaningfulness of 7-letter change in LCLA. The clinical meaningfulness of the 7-letter LCLA worsening has been demonstrated previously, using validated visual quality of life scales.⁹

Limitations of this work include the availability of somewhat fewer data for LCLA, SDMT, and self-reported measures. Relatively few datasets contained all 4 performance measures and EDSS, limiting the analyses of their relative sensitivity, including in RR vs progressive MS, as a function of baseline disability level, and to capture worsening disability independent of relapse. In addition, our ability to fully explore clinical meaningfulness of the performance measures using self-report measures also was restricted by lack of self-report measures used across studies beyond SF-36 and BDI. Other measures of self-report have been applied to the MS population, but these analyses were limited by the surveys in the existing dataset, that is, the SF-36. Also, although the dataset included the full range of disability, the majority of participants had RRMS and relatively mild disability, with median EDSS of 2.5, reflecting the over-representation of clinical trials in RRMS in the MS field. This point may limit a full understanding of the performance tests in more disabled, progressive populations. Finally, for these analyses, we pooled treatment groups and focused on 3-month confirmed disability worsening (rather than 6-month), which could have affected the results.

These results confirm the advantageous measurement properties of the T25FW, 9HPT, LCLA, and SDMT and support their construct, convergent, and known group validity, and sensitivity, particularly when combined into a multidimensional test battery. The associations with established measures of disability (EDSS) and HRQoL (SF-36) indicate that they evaluate clinically meaningful aspects of MS-related disability. These findings support the use of these measures either alone or together as a multidimensional test battery as primary or key secondary endpoints in MS studies.

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Appendix Authors

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Nicholas G. LaRocca, PhD	National Multiple Sclerosis Society, New York, NY	Author	Conceptualized and designed the study, interpreted the results, reviewed and edited the manuscript for intellectual content
Richard A. Rudick, MD	Biogen, Cambridge, MA	Author	Conceptualized and designed the study, interpreted the results, reviewed and edited the manuscript for intellectual content
Lynn D. Hudson, PhD	Critical Path Institute, Tucson, AZ	Author	Conceptualized and designed the study, interpreted the results, reviewed and edited the manuscript for intellectual content
Peter Chin, MD	Genentech, San Francisco, CA	Author	Conceptualized and designed the study, interpreted the results, reviewed and edited the manuscript for intellectual content
Gordon S. Francis, MD	At onset of MSOAC, employee of Novartis; currently independent neurology clinical development consultant, San Francisco, CA	Author	Conceptualized and designed the study, interpreted the results, reviewed and edited the manuscript for intellectual content
Adam Jacobs, PhD	Premier Research, Quincy, MA	Author	Conceptualized and designed the study, performed the statistical analyses, interpreted the results, reviewed and edited the manuscript for intellectual content
Raj Kapoor, FRCP	UCL Institute of Neurology, London, UK	Author	Conceptualized and designed the study, interpreted the results, reviewed and edited the manuscript for intellectual content
Paul M. Matthews, MD	Imperial College, London, UK	Author	Conceptualized and designed the study, interpreted the results, reviewed and edited the manuscript for intellectual content

Appendix (continued)

Name	Location	Role	Contribution
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References

- Whitaker JN, McFarland HF, Rudge P, Reingold SC. Outcomes assessment in multiple sclerosis clinical trials: a critical analysis. *Mult Scler* 1995;1:37–47.
- Rudick R, Antel J, Confavreux C, et al. Clinical outcomes assessment in multiple sclerosis. *Ann Neurol* 1996;40:469–479.
- Rudick R, Antel J, Confavreux C, et al. Recommendations from the National Multiple Sclerosis Society clinical outcomes assessment task force. *Ann Neurol* 1997;42:379–382.
- Schwid SR, Goodman AD, McDermott MP, Bever CF, Cook SD. Quantitative functional measures in MS: what is a reliable change? *Neurology* 2002;58:1294–1296.
- Kragt J, Van der Linden F, Nielsen J, Uitdehaag B, Polman C. Clinical impact of 20% worsening on Timed 25-Foot Walk and 9-Hole Peg Test in multiple sclerosis. *Mult Scler* 2006;12:594–598.
- van Winsen L, Kragt J, Hoogervorst E, Polman C, Uitdehaag B. Outcome measurement in multiple sclerosis: detection of clinically relevant improvement. *Mult Scler* 2010;16:604–610.
- Bosma L, Kragt J, Brieve L, et al. Progression on the Multiple Sclerosis Functional Composite in multiple sclerosis: what is the optimal cut-off for the three components? *Mult Scler* 2010;16:862–867.
- Balcer LJ, Miller DH, Reingold SC, Cohen JA. Vision and vision-related outcome measures in multiple sclerosis. *Brain* 2015;138:11–27.
- Balcer LJ, Raynowska J, Nolan R, et al. Validity of low-contrast letter acuity as a visual performance outcome measure for multiple sclerosis. *Mult Scler J* 2017;23:734–747.
- Benedict RHB, DeLuca J, Phillips G, et al. Validity of the Symbol Digit Modalities Test as a cognition performance outcome measure for multiple sclerosis. *Mult Scler J* 2017;23:721–733.
- Ontaneda D, LaRocca N, Coetzee T, Rudick RA. Revisiting the multiple sclerosis functional composite: proceedings from the National Multiple Sclerosis Society (NMSS) task force on clinical disability outcomes. *Mult Scler J* 2012;18:1074–1080.
- Rudick RA, LaRocca N, Hudson LD; MSOAC. Multiple Sclerosis Outcome Assessments Consortium: genesis and initial project plan. *Mult Scler J* 2014;20:12–17.
- LaRocca NG, Hudson LD, Rudick R, et al. The MSOAC approach to developing performance outcomes to measure and monitor multiple sclerosis disability. *Mult Scler J* 2018;24:1469–1484.
- Multiple Sclerosis Outcome Assessments Consortium and the CFAST Multiple Sclerosis Development Team. Therapeutic area data standards user guide for multiple sclerosis version 1.0 [online]. Available at: cdisc.org/therapeutic#MS. Accessed February 22, 2019.
- Motl RW, Cohen JA, Benedict R, et al. Validity of the timed 25-foot walk as an ambulatory performance outcome measure for multiple sclerosis. *Mult Scler J* 2017;23:704–710.
- Feys P, Lamers I, Francis G, et al. The Nine-Hole Peg Test as a manual dexterity performance measure for multiple sclerosis. *Mult Scler J* 2017;23:711–720.
- Strober L, DeLuca J, Benedict RHB, et al. Symbol Digit Modalities Test: a valid clinical trial endpoint for measuring cognition in multiple sclerosis. *Mult Scler J* 2018 Epub ahead of print.
- Kurtzke JF. Rating neurologic impairment in multiple sclerosis: an expanded disability status scale (EDSS). *Neurology* 1983;33:1444–1452.
- Ware JE, Snow KK, Kosinski M, Gandek B. SF-36 Health Survey: Manual and Interpretation Guide. Boston: The Health Institute, New England Medical Center; 1993.
- Beck AT, Ward CH, Mendelson M, Mock J, Erbaugh J. An inventory for measuring depression. *Arch Gen Psychiatry* 1961;4:561–571.
- Cohen JA, Reingold SC, Polman CH, Wolinsky JS; International Advisory Committee on Clinical Trials in Multiple Sclerosis. Disability outcome measures in multiple sclerosis trials: current status and future prospects. *Lancet Neurol* 2012;11:467–476.
- Norman GR, Sloan JA, Wyrwich KW. Interpretation of changes in health-related quality of life: the remarkable universality of half a standard deviation. *Med Care* 2003;41:582–592.
- Fischer JS, Jak AJ, Knicker JE, Rudick RA, Cutter G. Multiple Sclerosis Functional Composite (MSFC). Administration and Scoring Manual. Unitech: National Multiple Sclerosis Society; 2001.
- Winer BJ. *Statistical Principles in Experimental Design*. 2nd ed. New York: McGraw-Hill; 1971.
- Critical Path Institute. Available at: c-path.org. Accessed February 19, 2019.

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Evaluation of multiple sclerosis disability outcome measures using pooled clinical trial data

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Disputes & Debates: Editors' Choice

Steven Galetta, MD, FAAN, Section Editor

Editors' Note: Association of Prestroke Metformin Use, Stroke Severity, and Thrombolysis Outcome

In "Association of prestroke metformin use, stroke severity, and thrombolysis outcome," Westphal et al. reviewed data from the European Thrombolysis in Ischemic Stroke Patients collaboration and compared stroke severity (NIH Stroke Scale), 3-month functional outcome (modified Rankin score), and mortality in 757 patients with type 2 diabetes who received metformin before stroke and 1,162 patients with type 2 diabetes who did not receive metformin before stroke. The authors reported that patients on metformin had less severe strokes, better functional outcome, and lower mortality and concluded that these findings suggest metformin has a protective effect in this patient population. Although he found the findings interesting, Dasheiff believed that this conclusion was premature because (1) correlation does not imply causation, (2) the statistical analysis did not fully compensate for the myriad confounders, and (3) there is no mention of other diabetic medications such as insulin use, which, in addition to the HbA1c and blood glucose may reflect disease severity. Westphal et al. agreed that causality cannot be inferred but emphasized the strengths of their study including the large sample size and their use of propensity score matching. The authors suggested their findings may precipitate additional studies on the relationship between metformin and outcome after stroke and increased use of metformin in patients with diabetes and vascular risk factors.

Ariane Lewis, MD, and Steven Galetta, MD
Neurology® 2021;96:500. doi:10.1212/WNL.0000000000011565

Reader Response: Association of Prestroke Metformin Use, Stroke Severity, and Thrombolysis Outcome

Richard Dasheiff (Plano, TX)
Neurology® 2021;96:500–501. doi:10.1212/WNL.0000000000011564

When I first read the article by Westphal et al.,¹ I was excited about the possible positive effects of metformin on stroke. I note that a one-point difference in NIHSS is not clinically significant, a one-point change in modified Rankin Score is somewhat significant, and 13% vs 18% death rate is more clinically significant. However, further critical analysis challenges this interpretation. First, the authors conclude that "[t]his suggests a protective effect of MET [metformin]." *Neurology* reviewers and editors should admonish authors about inferring causality from inappropriately designed studies. In addition, as the authors report, "The number of confounders in this study was 19; together with the treatment group (MET+/MET-), 20 independent variables resulted." This can only partially be compensated by their propensity score matching and multiple imputation. I would point out several other problems. For example, missing values included international normalized ratio (INR) on admission (29.7%). How could this be if everyone got IV thrombolysis? Per the article by Scheitz et al. cited in reference 8, "Laboratory measures obtained on hospital admission include glucose (mmol/L), blood cell counts (platelets, hemoglobin, and leucocytes), international normalized ratio and partial thromboplastin time/ Not all centers have to provide data on all variables but have given a commitment to add missing variables retrospectively, if considered relevant to answer a specific research question."

Author disclosures are available upon request (journal@neurology.org).

So, another question is whether everyone in this study had an INR before IV thrombolysis? Finally, perhaps the most important factor is the severity of the subject's type 2 diabetes. Although blood glucose and HbA1c are surrogates for the patient's clinical state, there is no assessment of other diabetes comedications. Patients usually move along a spectrum of diet control, oral hypoglycemics, and insulin. The composition of the MET group could be weighted toward insulin dependence, which would make them, on average, treated longer for their diabetes and having a more severe illness. That alone might account for their worse clinical outcomes, and not be a protective effect of metformin.

1. Westphal LP, Widmer R, Held U, et al; Thrombolysis in Ischemic Stroke Patients (TRISP) Study Group. Association of prestroke metformin use, stroke severity, and thrombolysis outcome. *Neurology* 2020;95:e362–e373.

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Author Response: Association of Prestroke Metformin Use, Stroke Severity, and Thrombolysis Outcome

Laura P. Westphal (Zurich), Ulrike Held (Zurich), Stefan Engelter (Basel, Switzerland), and Susanne Wegener (Zurich)

Neurology® 2021;96:501. doi:10.1212/WNL.0000000000011563

We thank Dr. Dasheiff for his interest in our article.¹ We agree that based on observational registry-based data, it is not straightforward to claim causality and impossible to account for all, including unknown, confounders. However, a thorough analysis based on propensity score–matched (PSM) sets of patients is mimicking a randomized experiment, and because the data set was very large, we were able to account for many known potential confounders simultaneously. It is common in real-life databases that there are some observations missing. In our study, we tackled this with a complex approach, in which we performed imputation and PSM repeatedly in multiple iterative steps. The observation that international normalized ratio (INR) values were incomplete is explained by the fact that standard operating procedures for intravenous thrombolysis (IVT) differed across centers. Some centers do not determine INR before IVT in patients with neither anticoagulation nor liver or coagulation disorder. We agree that duration of diabetes, nonpharmacological and other cotreatments may matter. This limitation was mentioned in the study. By accounting for the age difference between metformin (MET)-/MET+ groups with PSM, we intended to address this aspect. Our data may serve as an argument for patients and GPs to continue MET in patients with diabetes and vascular risk factors and spur planning of further trials that may prove the suggested protective effect of MET.

1. Westphal LP, Widmer R, Held U, et al; Thrombolysis in Ischemic Stroke Patients (TRISP) Study Group. Association of prestroke metformin use, stroke severity, and thrombolysis outcome. *Neurology* 2020;95:e362–e373.

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Editors' Note: Association of Guideline Publication and Delays to Treatment in Pediatric Status Epilepticus

In "Association of guideline publication and delays to treatment in pediatric status epilepticus," Fernandez et al. reported that despite publication of evidence of delays in treatment of refractory status epilepticus (SE) in the Pediatric Status Epilepticus Research Group (pSERG) at the end of 2014, there was no difference in time to initiation of benzodiazepines, nonbenzodiazepine antiepileptic drugs (AEDs), and continuous infusions to patients with refractory SE in pSERG in the hospital between 2011–2014 and 2015–2019. The authors identified a number of proposed actions to overcome potential barriers to the timely initiation of AEDs. Albuja et al. commented that they improved the time to administration of second-line AEDs at their institution by creating a SE alert system to contact neurology, pharmacy, the rapid response team, and the bed manager simultaneously when an inpatient is suspected to be in SE. Amengual-Gual et al. applauded them for implementing this initiative. It is imperative to continue to identify quality improvement measures to identify SE and ensure that AEDs are given in a timely fashion to patients with SE because it becomes more resistant over time, which may lead to increased morbidity and mortality.^{1,2}

Ariane Lewis, MD, and Steven Galetta, MD
Neurology® 2021;96:502. doi:10.1212/WNL.0000000000011562

1. Sánchez Fernández I, Abend NS, Agadi S, et al. Gaps and opportunities in refractory status epilepticus research in children: a multi-center approach by the Pediatric Status Epilepticus Research Group (pSERG). *Seizure* 2014;23:87–97.
2. Madžar D, Geyer A, Knappe RU, et al. Association of seizure duration and outcome in refractory status epilepticus. *J Neurol* 2016;263:485–491.

Reader Response: Association of Guideline Publication and Delays to Treatment in Pediatric Status Epilepticus

Ana C. Albuja (Providence, RI), Meriem K. Bensalem-Owen (Lexington, KY), and
Mauricio F. Villamar (Providence, RI)
Neurology® 2021;96:502–503. doi:10.1212/WNL.0000000000011561

Sánchez Fernández et al.¹ reported that the publication of evidence on delays in treatment of pediatric refractory convulsive status epilepticus (rSE) was not associated with improvement in time to treatment of rSE (TTTSE). They propose interventions to reduce TTTSE. We would like to share a successful model that we implemented. In 2013, we identified delays in TTTSE in our institution.² We saw no improvement in TTTSE after a dual intervention consisting of educational programs for healthcare professionals and the development of an electronic SE order set.³ This prompted us to develop an SE alert system that was modeled after our "code stroke" protocol.⁴ When an inpatient is suspected to have SE, the SE alert is activated by calling an operator who simultaneously pages the neurology and pharmacy house staff, the rapid response team, and the manager responsible for bed assignment. Use of this alert system led to considerable reduction in time to administration of second-line antiseizure medications (22.21 ± 3.44 minutes) compared with standard care (58.30 ± 6.72 minutes; $p < 0.0001$).⁴ Because many hospitals have alert protocols for management of acute stroke, SE alert systems could be replicated with local resources to facilitate timely management and coordinated care of SE in adults and children.

1. Sánchez Fernández I, Abend NS, Amengual-Gual M, et al. Association of guideline publication and delays to treatment in pediatric status epilepticus. *Neurology* 2020;95:e1222–e1235.
2. Boske A, Cook A, Bensalem Owen M. Assessment of time of administration from time of order of anti-epileptic drugs in status epilepticus. *Neurology* 2013;80:S48.007.
3. Callow M, Cook A, Mahmoud A, et al. Assessment of the timeliness of administration of second line antiepileptic drugs for status epilepticus after the implementation of a status epilepticus bundle order set: a single institution experience. *Neurology* 2015;84:P7.322.
4. Villamar MF, Cook AM, Ke C, et al. Status epilepticus alert reduces time to administration of second-line antiseizure medications. *Neurol Clin Pract* 2018;8:486–491.

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Author Response: Association of Guideline Publication and Delays to Treatment in Pediatric Status Epilepticus

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Neurology® 2021;96:503. doi:10.1212/WNL.00000000000011560

We thank Drs. Albuja, Bensalem-Owen, and Villamar for their comment on our study¹ and for their crucial contribution, highlighting the developments toward improving status epilepticus treatment in this pilot study in 19 adult patients during inpatient care at a single center² and our publication. We are entirely aligned with these essential suggestions and applaud the authors for their crucial contribution. Similarly—and simultaneously to the authors' accomplishment—pediatric investigators were able to improve inpatient treatment times in children at a single center.³ Most ongoing delays continue to occur in patients with intermittent refractory status epilepticus—related to difficulties in assessment whether convulsive status epilepticus persists—and in the outpatient setting.⁴ Our group is actively working on measures that can identify ongoing seizures—not only at a single center but also for larger populations in multicenter settings and in the outpatient setting.⁵ We agree that—among other approaches such as implementing novel technologies and quality improvement techniques—learning from previous stroke treatment efforts will play an important role.⁶ We are grateful for the chance to add this outlook to our current article and congratulate Drs. Albuja, Bensalem-Owen, and Villamar on this innovative milestone toward our shared goal of implementing improved status epilepticus care for our patients.

1. Sánchez Fernández I, Abend NS, Amengual-Gual M, et al. Association of guideline publication and delays to treatment in pediatric status epilepticus. *Neurology* 2020;95:e1222–e1235.
2. Villamar MF, Cook AM, Ke C, et al. Status epilepticus alert reduces time to administration of second-line antiseizure medications. *Neurol Clin Pract* 2018;8:486–491.
3. Adam P, Ostendorf AP, Merison K, Wheeler TA, Patel AD. Decreasing seizure treatment time through quality improvement reduces critical care utilization. *Pediatr Neurol* 2018;85:58–66.
4. Sánchez Fernández I, Gaínza-Lein M, Abend NS, et al. Factors associated with treatment delays in pediatric refractory convulsive status epilepticus. *Neurology* 2018;90:e1692–e1701.
5. Gaínza-Lein M, Sánchez Fernández I, Ulate-Campos A, Loddenkemper T, Ostendorf A. Timing in the treatment of status epilepticus: from basics to the clinic. *Seizure* 2019;68:22–30.
6. Amengual-Gual M, Ulate-Campos A, Loddenkemper T. Status epilepticus prevention, ambulatory monitoring, early seizure detection and prediction in at-risk patients. *Seizure* 2019;68:31–37.

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Evaluation of Multiple Sclerosis Disability Outcome Measures Using Pooled Clinical Trial Data

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In the article “Evaluation of Multiple Sclerosis Disability Outcome Measures Using Pooled Clinical Trial Data” by M. Goldman et al.,¹ the text contained some discrepancies within the analyses of the MSOAC data set included in the paper. The authors found that the discrepancies were present because of a coding error that produced inaccurate results in the Kaplan-Meier results. The authors regret the errors. The text below, provided by the authors, describes the corrected material:

From the Authors:

In our publication in *Neurology*,¹ we analyzed a pooled dataset comprising 12,776 clinical trial participants that had been assembled by the Multiple Sclerosis Outcome Assessments Consortium (MSOAC) to evaluate 4 performance tests as proposed components of a multidimensional test battery. We reported measurement properties; construct, convergent, and known group validity; and longitudinal performance of the Timed 25-Foot Walk (T25FW), 9-Hole Peg Test (9HPT), Low Contrast Letter Acuity (LCLA), and Symbol Digit Modalities Test (SDMT) individually and when combined into a multidimensional test battery compared to the Expanded Disability Status Scale (EDSS) and Short-Form-36 Physical Component Summary. The placebo arm data in the MSOAC database were made publicly available to support research by investigators in the MS field. Following publication of our paper, a colleague, using the publicly available placebo data from MSOAC, contacted us after finding seemingly different results than reported in our paper. To better understand the potential discrepancy in results, we returned to our analysis. Subsequently, errors in statistical program (Statistical Analysis System, SAS) coding for some of our analyses were discovered. These errors resulted in incorrect progression rates for the T25FW and 9HPT, as well as the composite measures of progression based on any 1 or any 2 measures, which included T25FW and 9HPT. In addition, while the graph of progression by LCLA was correct, in the text, we incorrectly gave results for LCLA progression (and its kappa coefficient for agreement with EDSS) using a 20% threshold, instead of the 7-point threshold stated. Herein, we report corrected results.

Table 1 Sensitivity to Change of Performance Tests and Expand Disability Status Scale

Performance test	Kaplan-Meier estimates of the percentage (95% CI) of trial participants with 3-mo confirmed disability progression over 24 mo		
	Previously reported	Corrected	EDSS ^{e,f}
T25FW ^a	6.5	23.8 (22.8–24.8)	20.2 (19.2–21.1)
9HPT ^b	2.9	8.6 (8.0–9.3)	20.2 (19.2–21.1)
LCLA ^c	13.1	13.8 (12.7–14.9)	16.1 (15.1–17.2)
SDMT ^d	15.0 ^g	15.0 ^g (13.4–16.7)	14.5 (12.9–16.2)
Worsening on ≥1 performance test	Not previously reported	41.5 (38.9–44.1)	11.5 (9.9–13.3)
Worsening on ≥2 performance tests	Not previously reported	12.9 (11.2–14.8)	11.5 (9.9–13.3)

Abbreviations: 9HPT = 9-Hole Peg Test; CI = confidence interval; EDSS = Expanded Disability Status Scale; LCLA = Low-Contrast Letter Acuity; SDMT = Symbol Digit Modalities Test; T25FW = Timed 25-Foot Walk.

^a 20% threshold.

^b 20% threshold.

^c Seven-point threshold: the previous paper had erroneously reported results for a 20% threshold.

^d Four-point threshold.

^e Baseline score 0: 1.5-point increase, baseline score 1.0–5.5: 1-point increase, baseline score ≥6.0: 0.5-point increase.

^f The study populations available for each comparison differed, leading to differing proportions with 3-mo confirmed worsening on EDSS.

^g The previously reported result for 3-mo confirmed worsening over 18 mo was correct. For consistency, we report the rate over 24 mo here.

Table 2 Associations of 3-Month Confirmed Progression Events Defined by Performance Tests and by Expanded Disability Status Scale

Comparison	Cohen's κ (95% CI) Previously reported	Corrected
T25FW ^a vs EDSS ^e	0.02 (−0.00 to 0.03)	0.37 (0.35 to 0.39)
9HPT ^b vs EDSS ^e	0.00 (−0.01 to 0.01)	0.25 (0.22 to 0.27)
LCLA ^c vs EDSS ^e	0.11 (0.08 to 0.14) ^f	0.10 (0.07 to 0.13)
SDMT ^d vs EDSS ^e	−0.02 (−0.06 to 0.02)	−0.02 (−0.06 to 0.02)

Abbreviations: 9HPT = 9-Hole Peg Test; CI = confidence interval; EDSS = Expanded Disability Status Scale; LCLA = Low-Contrast Letter Acuity; SDMT = Symbol Digit Modalities Test; T25FW = Timed 25-Foot Walk.

^a 20% threshold.

^b 20% threshold.

^c Seven-point threshold (results had previously been given for a 20% threshold).

^d Four-point threshold.

^e Baseline score 0: 1.5-point increase, baseline score 1.0–5.5: 1-point increase, baseline score ≥ 6.0 : 0.5-point increase.

^f The previously reported association between progression events defined by 20% worsening on LCLA and EDSS was correct. For consistency, the association between progression events defined by 7-point worsening on LCLA and EDSS is reported here.

To compare the sensitivity to change of the performance measures with EDSS, time from baseline to 3-month confirmed worsening over 24 months was analyzed (table 1). We previously reported that the progression rates were lower for T25FW and 9HPT compared to EDSS. In the corrected analyses, the progression rate remained somewhat lower for 9HPT compared to EDSS, while progression rates for T25FW, LCLA, and SDMT were similar to or higher than the EDSS (results for LCLA and SDMT were the same as before, as the programming error affected only T25FW and 9HPT). When the performance tests were combined into a multidimensional outcome measure, the proportion of participants with confirmed worsening on any 1 performance test was substantially greater than the proportion with confirmed worsening on EDSS. When confirmed worsening on 2 performance tests was required, sensitivity to disability progression was similar to that of EDSS. Association of the progression events defined by the performance tests was better correlated with those defined by the EDSS in the corrected analysis, though remained modest (table 2).

Based on the previously reported analyses, we concluded that the results supported the use of the T25FW, 9HPT, LCLA, and SDMT as study outcome measures, both individually or combined into a multidimensional test battery.¹ These corrected results demonstrate better sensitivity to change for the T25FW and 9HPT than previously reported, and further support our original conclusion.

Reference

1. Goldman MD, LaRocca N, Rudick R, et al. Evaluation of multiple sclerosis disability outcome measures using pooled clinical trial data. *Neurology* 2019;93:e1921–e1931.